# On the moments of a distribution defined by the Gaussian polynomials 

Helmut Prodinger<br>School of Mathematics, The John Knopfmacher Centre for Applicable Analysis and Number Theory, University of the Witwatersrand, P.O. Wits, 2050 Johannesburg, South Africa

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#### Abstract

An alternative method is presented to compute the moments of the probability distribution defined by the Gaussian polynomials. It computes the cumulants first. (c) 2002 Elsevier B.V. All rights reserved.


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Di Bucchianico (1999) has considered the probability distribution defined by the (probability) generating function

$$
F_{m, n}(q)=\frac{\left[\begin{array}{c}
m+n \\
m
\end{array}\right]_{q}}{\binom{m+n}{m}}
$$

with the Gaussian polynomials

$$
\left[\begin{array}{c}
m+n \\
m
\end{array}\right]_{q}=\frac{(1-q)\left(1-q^{2}\right) \cdots\left(1-q^{m+n}\right)}{(1-q)\left(1-q^{2}\right) \cdots\left(1-q^{m}\right)(1-q)\left(1-q^{2}\right) \cdots\left(1-q^{n}\right)}
$$

He discussed several methods to compute the moments of this distribution.
In this short note, I want to draw the attention of the reader to another (potentially superior) method that is due to Panny (1986). This method first computes the cumиlants and translates them into the moments by a standard formula. This yields explicit formulæ for all the moments.

[^0]Note that

$$
F_{m, n}(q)=\frac{g_{1}(q) \cdots g_{m+n}(q)}{g_{1}(q) \cdots g_{m}(q) g_{1}(q) \cdots g_{n}(q)}
$$

with

$$
g_{k}(q)=\frac{1-q^{k}}{k(1-q)}
$$

where $g_{k}(q)$ is a probability generating function of a random variable $X_{k}$. The function $\log g_{k}\left(\mathrm{e}^{\mathrm{i} t}\right)$ is the generating function of the cumulants. ${ }^{1}$ Panny has computed that as

$$
\log g_{k}\left(\mathrm{e}^{\mathrm{i} t}\right)=\frac{(k-1) \mathrm{i} t}{2}+\sum_{j \geqslant 1} \frac{B_{2 j}}{2 j}\left(k^{2 j}-1\right) \frac{(\mathrm{i} t)^{2 j}}{(2 j)!}, \quad|t|<\frac{2 \pi}{k},
$$

where $B_{i}$ denotes the $i$ th Bernoulli number.
Reading off coefficients we find the cumulants of $X_{k}$ as

$$
\kappa_{1}=\frac{k-1}{2}, \quad \kappa_{2 r}=\frac{\left(k^{2 r}-1\right) B_{2 r}}{2 r}, \quad \kappa_{2 r+1}=0, r=1,2,3, \ldots
$$

Since

$$
\log F_{m, n}\left(\mathrm{e}^{\mathrm{i} t}\right)=\sum_{k=1}^{m+n} \log g_{k}\left(\mathrm{e}^{\mathrm{i} t}\right)-\sum_{k=1}^{m} \log g_{k}\left(\mathrm{e}^{\mathrm{i} t}\right)-\sum_{k=1}^{n} \log g_{k}\left(\mathrm{e}^{\mathrm{i} t}\right),
$$

the cumulants $\kappa_{r}^{m, n}$ are obtained by summing the $\kappa_{r}$ 's.
So we get

$$
\begin{aligned}
\kappa_{1}^{m, n} & =\frac{\binom{m+n}{2}}{2}-\frac{\binom{m}{2}}{2}-\frac{\binom{n}{2}}{2}=\frac{m n}{2}, \\
\kappa_{2 r}^{m, n} & =\frac{B_{2 r}}{2 r(2 r+1)}\left(B_{2 r+1}(m+n+1)-B_{2 r+1}(m+1)-B_{2 r+1}(n+1)\right), \\
\kappa_{2 r+1}^{m, n} & =0
\end{aligned}
$$

Here is a little list:

$$
\begin{aligned}
\kappa_{1}^{m, n} & =\frac{m n}{2} \\
\kappa_{2}^{m, n} & =\frac{m n(m+n+1)}{12} \\
\kappa_{4}^{m, n} & =-\frac{m n(m+n+1)(m(m+1)+m n+n(n+1))}{120} .
\end{aligned}
$$

[^1]The formula to compute the moments $\mu_{r}$ from this is

$$
\mu_{r}=\sum_{\pi_{1}+2 \pi_{2}+\cdots+r \pi_{r}=r}\left(\frac{\kappa_{1}}{1!}\right)^{\pi_{1}}\left(\frac{\kappa_{2}}{2!}\right)^{\pi_{2}} \cdots\left(\frac{\kappa_{r}}{r!}\right)^{\pi_{r}} \frac{r!}{\pi_{1}!\pi_{2}!\ldots \pi_{r}!},
$$

which is a finite sum since $\pi_{k} \in\{0,1, \ldots, r\}$. Hence

$$
\begin{aligned}
& \mu_{1}=\frac{m n}{2} \\
& \mu_{2}=\kappa_{1}^{2}+\kappa_{2}=\frac{m n(m+n+3 m n+1)}{12} \\
& \mu_{3}=\kappa_{1}^{3}+3 \kappa_{1} \kappa_{2}+\kappa_{3}=\frac{m^{2}(m+1) n^{2}(n+1)}{8} \\
& \vdots
\end{aligned}
$$

We finish by mentioning that, as in (Panny, 1986), one could also get asymptotic formulæ.

## References

Di Bucchianico, A., 1999. Combinatorics, computer algebra, and the Wilcoxon-Mann-Whitney test. J. Statist. Plann. Inference 79, 349-364.
Panny, W., 1986. A note on the higher moments of the expected behavior of straight insertion sort. Inform. Process. Lett. 22, 175-177.


[^0]:    E-mail address: helmut@maths.wits.ac.za (H. Prodinger).
    URL: http://www.wits.ac.za/helmut/index.htm

[^1]:    ${ }^{1}$ We write $\log x$ for the (natural) logarithm with base $e$.

